








## RESEARCH ARTICLE

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# Extreme Event Statistics in Dst, SYM-H, and SMR Geomagnetic Indices

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### Key Points:

- These indices are not totally interchangeable, consideration should be given to index choice in model validation or cross-study comparison
- Hourly averaged SMR and SYM-H return levels track Dst for return periods below 10 years. Above that they exceed Dst; at 100 years by >10%
- One minute cadence SMR and SYM-H 5, 10, 50, and 100 year return levels exceed that of Dst by about 10%, 12%, 20%, and 25% respectively

### Supporting Information:

Supporting Information may be found in the online version of this article.

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**Abstract** Extreme space weather events are rare, and quantifying their likelihood is challenging, often relying on geomagnetic indices obtained from ground-based magnetometer observations that span multiple solar cycles. The Dst index ring-current monitor, derived from an hourly average over four low-latitude stations, is a benchmark for extreme space weather events, and has been extensively studied statistically. We apply extreme value theory (EVT) to two geomagnetic ring current indices: SYM-H (derived from 6 stations) and SMR (derived from up to 120 stations). EVT analysis reveals a divergence between the return level found for Dst, and those for SYM-H and SMR, that increases non-linearly with return period. For return periods below 10 years, hourly averaged SYM-H and SMR have return levels similar to Dst, but at return periods of 50 and 100 years, they respectively exceed that of Dst by about 10% and 15% (SYM-H) and about 7% and 12% (SMR). One minute resolution SYM-H and SMR return levels progressively exceed that of Dst; their 5, 10, 50, and 100 year return levels exceed that of Dst by about 10%, 12%, 20% and 25% respectively. Our results indicate that consideration should be given to the differences between the indices if selecting one to use as a benchmark in model validation or resilience planning for the wide range of space weather sensitive systems that underpin our society.

**Plain Language Summary** Extreme space weather events, which disturb Earth's near-Earth plasma and magnetic environment, have the potential to cause significant disruption to our infrastructure. Extreme events are rare, and quantifying their likelihood relies upon long-term continuous observations. High-quality ground-based magnetometer observations underpin geomagnetic indices that monitor space weather. The Dst index has become the de-facto measure for space weather storms. Space weather storms cause magnetic perturbation that can be localized in space and time. We now have the SYM-H and SuperMAG SMR indices, constructed to be 1-min versions of Dst, with SMR using a larger set of magnetometers. We perform the first extreme value theory analysis of SYM-H and SMR. We find that space weather storms are more intense when measured by the 1-min SYM-H or SMR indices but the relationship between extreme values identified in the Dst timeseries and those in SYM-H or SMR is not linear. Our results indicate that consideration should be given to the differences between the indices if selecting one to use as a benchmark in model validation or resilience planning for the wide range of space weather sensitive systems that underpin our society.

## 1. Introduction

Severe geomagnetic storms can give rise to widespread societal, economic, and technological disruption (Hapgood et al., 2021). Rare events, such as the 1859 Carrington storm (Cliver & Dietrich, 2013), and the 1921 storm (Hapgood, 2019) have been estimated to potentially result in nationwide disruption of power distribution (Oughton et al., 2017) or significant satellite loss (Horne et al., 2013). Consequently, long-term forecasting of magnetospheric space weather is of significant operational interest (Morley, 2020). A key element of an extreme space weather event is a dramatic and rapid enhancement of the large scale currents carried by the ionosphere and magnetosphere. This perturbs the magnetic field at the surface of the earth. Ground-based magnetometer observations, spanning multiple solar cycles, form the basis of geomagnetic indices which are routinely used

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to quantify magnetospheric activity (Mayaud, 1980). Geomagnetic storms drive enhancements of earth's ring current and this can be monitored by low- and mid-latitude magnetometer stations.

The hourly Dst index is based upon an average of observations from four such low-latitude stations (Sugiura, 1964), and is commonly used in the characterization of geomagnetic storms (Gonzalez et al., 1994). Whilst it is understood that the detailed spatial and temporal evolution of a storm, which map directly onto its detailed impact, are not consistently captured by ring current indices (e.g., Sandhu et al. (2021) and references therein), the Dst index combines fidelity with multi-solar cycle coverage. Indices built on ground based magnetometers will inevitably sample space weather events in an inhomogeneous manner. The location of the ground based stations relative to the excited storm-time current systems will vary from one event to another, as will the number of stations and their duty cycles. Our results will therefore provide a lower bound on the event return level, which we anticipate will increase with improved spatial and temporal sampling in SMR as compared to Dst, particularly if these indices respond to current systems other than the symmetric ring current.

Storm occurrence rates and intensities vary with the approximately 11 years cycle of solar activity (Bergin et al., 2022; Chapman, McIntosh, et al., 2020) and each cycle is unique in amplitude and duration. Statistical studies to quantify space weather risk thus require observations spanning multiple solar cycles. Extreme events are by definition rare events, so that, whilst it is well known that their occurrence has a solar cycle modulation (Chapman, McIntosh, et al., 2020; Love, 2021), “benchmark” statistical estimates are based on observations across multiple cycles (Love et al., 2015; Riley, 2012). Provided the asymptotic tail of the underlying distribution (aggregated over a number of solar cycles) is sufficiently sampled, an extreme value analysis via peaks over threshold is still valid. As with previous work (Acero et al., 2018; Love et al., 2015; Riley, 2012; Tsubouchi & Omura, 2007), the study presented here should be understood as a solar cycle aggregate.

The Dst index is used as the de-facto benchmark in studies of space weather hazards (Cliver et al., 2022) and remains valuable as an indicator of geomagnetic field disturbance activity. The observed distribution of geomagnetic storm amplitudes, as seen in the maximum excursion of indices such as Dst, is routinely modeled to infer space weather risk. Such modeling provides an estimate of the amplitude of rare, extreme, “one in a hundred year” events. From physical constraints, there will be an upper limit to the size of such events, estimated at  $Dst \sim -2,000$  nT to  $-2,500$  nT (Liu et al., 2020; Vasyliūnas, 2011). Events that have occurred in the past 100 years fall within this limit. For example, the 1921 storm which today could cause continent-wide disruption (Hapgood, 2019; Oughton et al., 2017) had a Dst estimated at  $\sim -900$  nT, comparable with the Carrington event (Love et al., 2019); whereas the 1989 storm which caused local disruption including a 9 hr power blackout in Quebec reached Dst of  $-589$  nT.

One approach to space weather risk inference is to fit different specific functions to the observed distributions of storm sizes (Love, 2020; Love et al., 2015; Riley, 2012). In the distribution tail this leads to a wide range of estimates. For example, for a power-law tail the probability of another extreme geomagnetic event comparable to the Carrington event occurring within the next 10 years is 10.3% 95% confidence interval (CI) [0.9,18.7] but is only 3.0% 95% CI[0.6,9.0] for a lognormal distribution (Riley & Love, 2017). Confidence intervals are generally large (Love et al., 2015; Riley, 2012). To increase the number of events in the sample and to cover a wider range of solar activity, lower fidelity indices that exist for longer time intervals have been re-engineered to estimate Dst, giving estimates for the 100-year return level of  $-663$  nT, with a bootstrap 68% CI[ $-694, -497$ ] nT (Love, 2021) based on the constrained Weibull model, and the 150-year Dst return level to be  $-809$  95% CI[ $-955, -663$ ] nT (Chapman, Horne, & Watkins, 2020) based on an exponential model. Estimated probabilities are sensitive to the assumed source distribution (Love, 2021; Riley & Love, 2017).

An alternative approach is extreme value theory (EVT) (see e.g., Coles, 2001; Embrechts et al., 1997; Leadbetter et al., 1983). EVT offers a systematic method for estimating the functional form of the source distribution, suitable for classes of distribution tail that encompass power law, log-normal and exponential. Previous studies have applied EVT to long-time record but relatively low fidelity indices (Elvidge, 2020; Koons, 2001; Silbergleit, 1999; Siscoe, 1976). The underlying observations are highly coarse grained compared to Dst (Chapman, Horne, & Watkins, 2020, and refs therein.). EVT was first applied to the Dst index for the years 1957–2001 (Tsubouchi & Omura, 2007), giving a 1 in 100 years storm level of  $-645.3$  95% CI[ $-754.5, -536.1$ ] nT, latterly for Dst over 1957–2014 this becomes  $-576.9$  95% CI[ $-680.44, -473.37$ ] nT as it incorporates more recent, quieter solar cycles.

In this work we will apply EVT to the 1-min resolution SMR and SYM-H indices, which have been constructed to mimic Dst but at considerably improved resolution. Ground based magnetometers typically have 1 min time resolution, and over three hundred are now operating globally. This data set is now being used through the SuperMAG collaboration (Gjerloev, 2012) to construct the SMR index which incorporates data from up to 120 mid and low latitude stations, and is produced at 1-min cadence (Newell & Gjerloev, 2012). With a greater station density and temporal resolution, SMR is more likely to resolve the space and time localized variations in storm time magnetospheric-ionospheric disturbances that are not consistently resolved by the Dst index (Newell & Gjerloev, 2012; Wanliss & Showalter, 2006). The SYM-H index is also produced at 1-min cadence, based on data from 6 stations from a network of 11 (Iyemori et al., 2010). The presence of more information content in SYM-H, compared to hourly Dst, has been reported (Bej et al., 2020). Although low-latitude indices such as Dst were designed to monitor the ring current, for more extreme events, other current systems can come into play, and can contribute to localized and intense disturbances. Regardless of their source, these ground disturbances and their temporal variations are potential space weather hazards.

We for the first time apply EVT to 1-min SMR and SYM-H indices. We isolate individual extreme events using the standard peak-over-threshold (POT) method and compare Dst events with those identified in 1-min cadence SYM-H and SMR, and their hourly averages. When the return level and period are empirically estimated for individual POT events, SYM-H and SMR track each other. For return periods less than about 10 years, Dst POT events track those of hourly averages SMR and SYM-H, and are systematically offset from POT events in 1 min SYM-H and SMR. The EVT analysis of the full set of POT events reveals a divergence between the return level found for Dst, and those for SMR and SYM-H, that increases non-linearly with return period.

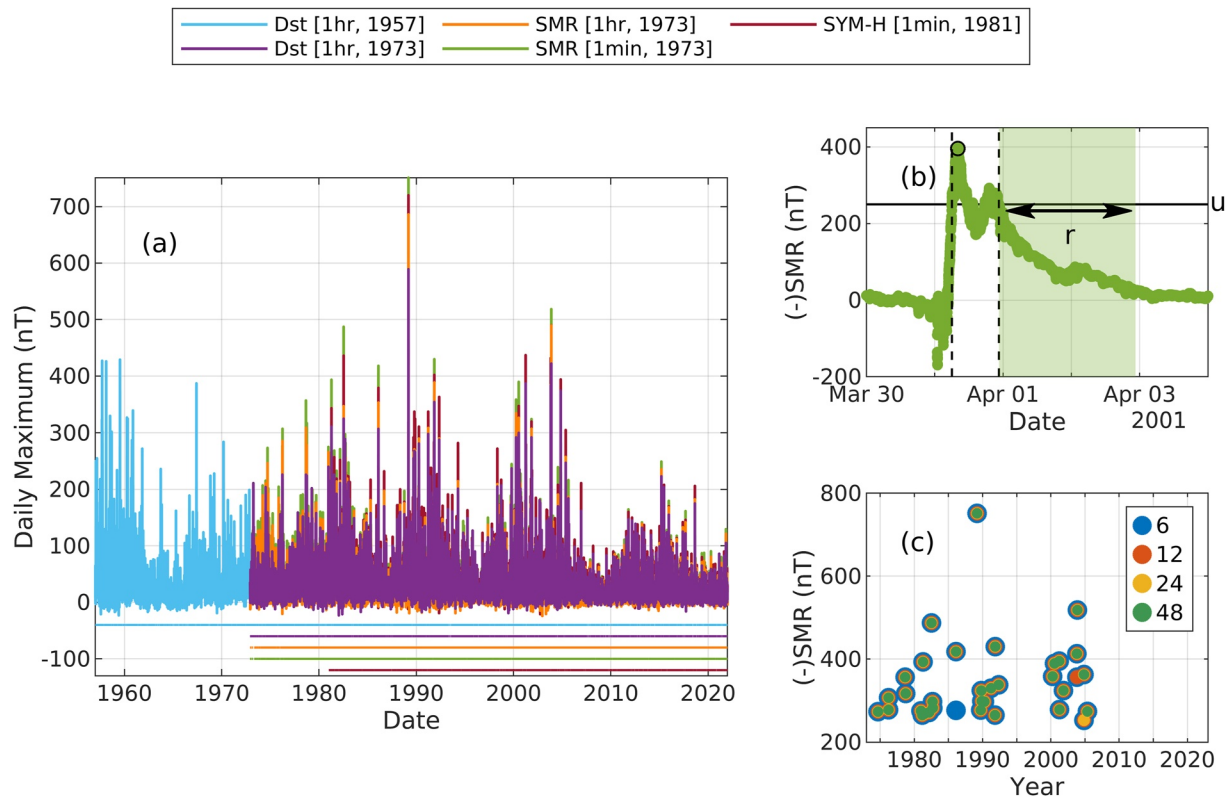
## 2. Methods

Three geomagnetic indices are of relevance for the present study; Dst, SYM-H, and SMR. The Dst index is compiled by the WDC for Geomagnetism Kyoto (Nose et al., 2015) and is available at 1-hr cadence. Four magnetometer stations located at geomagnetic latitudes between  $-35^{\circ}$  and  $+30^{\circ}$  record horizontal component disturbances in the geomagnetic field. Dst is then the baseline-corrected, magnetic latitude-adjusted, average of the four hourly mean disturbances. Dst for years 1957–2016 is termed “final” Dst and is “provisional” for years 2017–2021. The SYM-H index is also compiled by WDC for Geomagnetism Kyoto (Iyemori et al., 2010), it is available at 1-min cadence. Each month, six stations from the network of 11 geomagnetic observatories with geomagnetic latitudes between  $-50^{\circ}$  and  $+50^{\circ}$  are selected for the derivation of SYM-H, depending on data availability and quality. SYM-H is the baseline-corrected, magnetic latitude-adjusted, average of the disturbance component at each minute for the 6 stations. The SMR index (Newell & Gjerloev, 2012) is a SuperMAG index, based on all ground magnetometer stations contributing to the SuperMAG network at geomagnetic latitudes between  $-50^{\circ}$  and  $+50^{\circ}$ . The number of available stations has increased greatly over time, from the first available SMR observation based on less than 5 stations, to a present maximum count of 120 stations. The 1-min averages of baseline-corrected, magnetic latitude-adjusted, horizontal component disturbances are averaged within four equally sized local time sectors to produce four local time SMR indices, SMR is then the average of these four local time indices.

The following procedure is used in this study to select distinct extreme events: (a) exceedances of threshold,  $u$  are identified in the timeseries; (b) exceedances of  $u$  which are separated by a duration less than cluster interval,  $r$ , are labeled part of the same event, but when the time interval between successive exceedances is greater than  $r$  they constitute distinct events; (c) the peak (maximum) level recorded during each event is identified; (d) each event is represented by the timestamp and magnitude of the event peak. Following Acero et al. (2018) and Tsubouchi and Omura (2007), a cluster interval  $r = 48$  hr is used in the runs declustering of the geomagnetic indices throughout.

## 3. Results

We apply EVT to six geomagnetic index timeseries: the 1-min resolution SMR, available from 1973 (Newell & Gjerloev, 2012), 1-hr resolution Dst from the same interval, the full 1-hr resolution Dst, available from 1957 (Sugiura, 1964), an hourly average of SMR, which has the same temporal resolution as Dst, but is still constructed from a larger number of stations, the 1-min resolution SYM-H, available from 1981, and an hourly average of the SYM-H record. For convenience we will use minus the value of the index, for example,  $(-)$ Dst, throughout

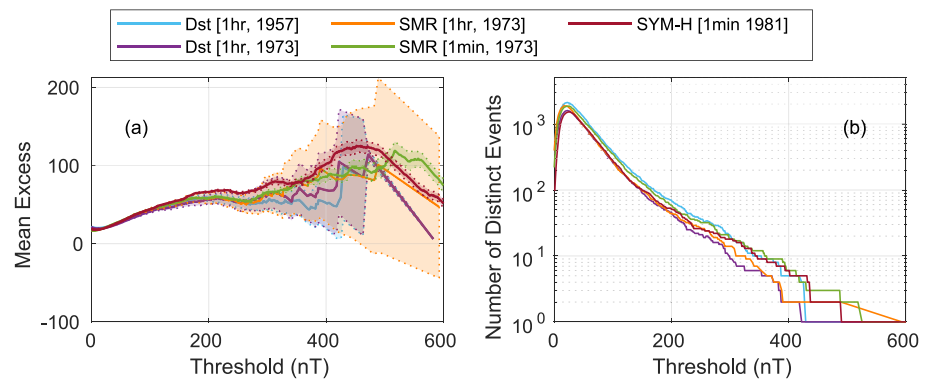


**Figure 1.** Data sets and declustering. (a) Daily maxima of indices are plotted with data availability indicated below the time series (horizontal line) for (-)Dst [1957–2021] (blue), (-)Dst [1973–2021] (purple), hourly average (-)SMR (orange), 1-min (-)SMR (green), and (-)SYM-H (red) indices. (b) Observations are plotted for (-)SMR [1-min, 1973] time series (green circles) between 30 March and 4 April 2001 showing how declustering is performed; threshold  $u = 250$  nT (black solid line), event start and end points (black dashed lines), cluster interval  $r = 48$  hr (green shading), and event peak-over-threshold (POT) (black open circle). (c) Event peak-over-thresholds for the (-)SMR [1-min, 1973] time series are plotted in time where distinct events are identified using cluster intervals  $r = 6$  hr (blue), 12 hr (orange), 24 hr (yellow), and 48 hr (green).

(neglecting the—sign, geomagnetic storm perturbations are negative). The timeseries are overplotted in Figure 1a. Activity can be seen to vary with the approximately 11 years solar cycle, and with cycle strength, it is relatively weak over the last 11 years (the weak cycle 24). Working with the SMR index from 1973 then misses the relatively active period (1957–1973), however the plot shows that the interval from 1973 does sample a wide range of activity levels. We present results for both the full Dst from 1957, and the Dst from 1973, contemporaneous with SMR. In Figure 1a we can also see that for the most extreme events, hourly averaged SMR often exceeds hourly Dst, and hourly indices are systematically exceeded by 1-min SMR and 1-min SYM-H. Improved spatial resolution (in SMR), and temporal resolution (in both SYM-H and SMR) may translate into increased estimates of return levels for extreme space weather events (Newell & Gjerloev, 2012; Wanliss & Showalter, 2006).

We need to identify individual extreme events in these timeseries. We will use the standard POT method (Davison & Smith, 1990) such that events are identified where observations systematically exceed some high threshold,  $u$ . EVT requires that the extreme events to be modeled are independent but, in geomagnetic index time series, temporal clustering of threshold exceedances is expected. Declustering is the procedure most commonly used to obtain a sample of independent events (Coles, 2001; Embrechts et al., 1997; Fawcett & Walshaw, 2012). Here, runs declustering (Smith, 1989) is applied to the timeseries, the result of which is shown in Figure 1b. We specify a minimum cluster interval  $r$  such that any sequence of  $\geq r$  consecutive observations below the threshold  $u$  discriminates two independent events. Following Acero et al. (2018) and Tsubouchi and Omura (2007), we use a cluster interval  $r = 48$  hr Figure 1c shows that for  $r > 6$  hr the results are insensitive to this choice.

EVT provides a framework to model the probability of extreme events that are rarely observed in a timeseries and to extrapolate these models to unobserved levels. The basis for the POT method (Davison & Smith, 1990) is that, for a stationary process, peak exceedances over a suitably high threshold, separated in time by a suitable cluster



**Figure 2.** Mean Excess and event histograms. (–)Dst [1957–2021] (blue), (–)Dst [1973–2021] (purple), hourly average (–) SMR (orange), 1-min (–)SMR (green), and (–)SYM-H (red) indices. (a) Mean excess plot (solid) with approximate 95% CI (shaded). (b) Histogram of events.

interval, have a limiting distribution which asymptotically approaches the generalized Pareto distribution (GPD) with survival function:

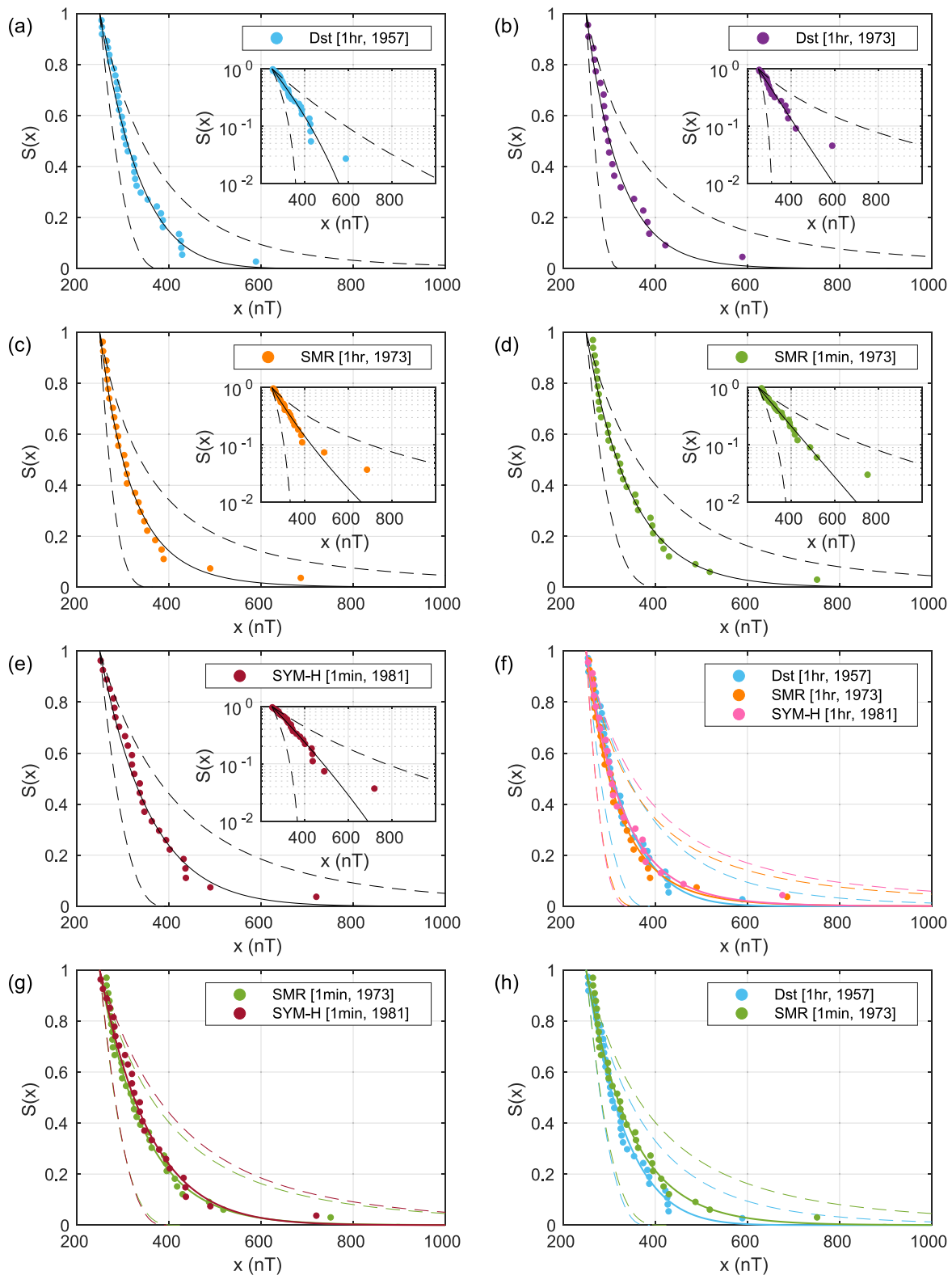
$$\bar{H}(y; \xi, \sigma) = \begin{cases} \left(1 + \frac{\xi y}{\sigma}\right)^{-1/\xi}, & \xi \neq 0 \\ \exp\left(-\frac{y}{\sigma}\right), & \xi = 0, \end{cases} \quad (1)$$

defined on  $\{y : y > 0 \text{ and } (1 + (\xi y)/\sigma) > 0\}$ , where the GPD limiting distribution has two parameters: scale ( $\sigma$ ), and shape ( $\xi$ ). We obtain the GPD for each geomagnetic index timeseries. The specification of an appropriately high threshold, here  $u = 250$  nT, for the GPD to be valid is addressed in the Supporting Information S1. A first indication of the GPD behavior is provided by a plot of the mean excess:

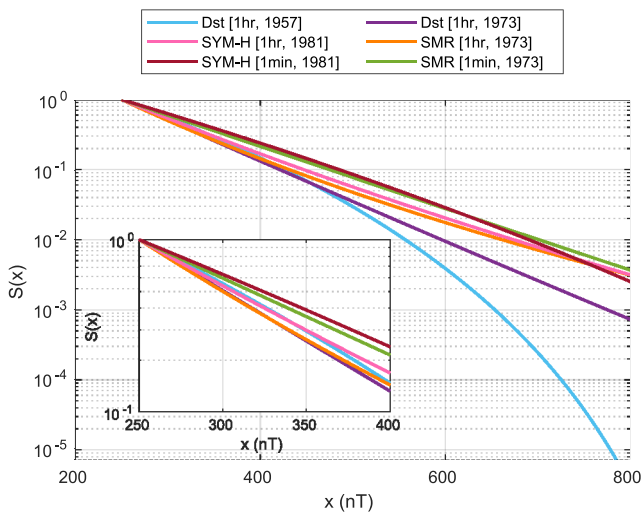
$$E(X - u | X > u) = \frac{1}{n_u} \sum_{i=1}^{n_u} (x_i - u) \quad (2)$$

where  $x_1, \dots, x_{n_u}$  are the  $n_u$  observations that exceed  $u$ . If a GPD is a reasonable model for excesses of a threshold  $u_0$ , for  $u > u_0$ ,  $E(X - u | X > u)$  is a linear function of  $u$ , with a gradient that determines the shape parameter. This is plotted in Figure 2a where we see that below 300 nT the mean excess is indistinguishable for Dst and SMR, but it diverges at higher thresholds. The SYM-H index mean excess systematically exceeds that of Dst or SMR samples over thresholds in the range 100–500 nT, but the shape of the SYM-H mean excess plot approximately tracks that of 1-min SMR. For a given time interval, the indices record approximately the same number of events, but 1-min indices have 60 times more data points than the 1-hr indices, so that the mean excess uncertainties are much smaller for 1-min SMR and SYM-H than 1-hr Dst or SMR timeseries. From Figure 2a, one can then discern for example, that the 1 min SMR is systematically more long-tailed than Dst; it has a larger positive gradient, hence larger positive shape ( $\xi$ ) parameter on the mean excess plot, suggesting more long-tailed behavior. Figure 2b plots histograms of the number of events identified in each timeseries as a function of declustering threshold. The longer interval, full Dst record contains more geomagnetic storms than the other timeseries, so can more frequently populate the far-tail. For data taken over the same time interval (from 1973), 1-hr average SMR matches, and can exceed, Dst. Both Dst and 1-hr SMR from 1973 are systematically exceeded by the sizes of events in 1-min SMR and SYM-H. For thresholds greater than 375 nT, the number of POT events in 1-min SYM-H (from 1981) exceeds that of Dst (from 1957), due to the higher values exhibited by the 1-min SYM-H index. For thresholds beyond 300 nT, the number of events in 1-min SMR (from 1973), exceeds that of Dst (from 1957).

We obtain GPD fits to these samples of events. The GPD model fit to the peaks-over-thresholds are shown in Figures 3 and 4 and the parameter estimates with their 95% confidence intervals are given in Table 1. The estimated GPD shape and scale parameters agree with each other; estimates for  $\xi$  are close to zero but the confidence intervals span the three families of the GPD;  $\xi < 0$ ,  $\xi = 0$ , and  $\xi > 0$ . Estimates of  $\sigma$  have large, overlapping confidence intervals. The more extreme events are, with the exception of the single largest event, reasonably



**Figure 3.** Generalized Pareto Distribution (GPD) models. Event peak-over-threshold values (filled circles) and corresponding GPD fit (black solid line) with 95% CI (black dashed line) for indices (a) (–)Dst [1957–2021] (blue), (b) (–)Dst [1973–2021] (purple), (c) hourly average (–)SMR (orange), (d) 1-min (–)SMR (green), and (e) 1-min (–)SYM-H (red). Insets repeat main plots with semi-log scale. (f) Overplot GPD fits and confidence intervals from (a) and (c) with GPD fit and 95% CI for hourly average (–)SYM-H index (pink). (g) Overplot GPD fits and confidence intervals from (d) and (e). (h) Overplot GPD fits and confidence intervals from (a) and (d).



**Figure 4.** Compare Generalized Pareto Distribution (GPD) models. GPD fits to event peak-over-threshold values for indices (–)Dst [1957–2021] (blue), (–)Dst [1973–2021] (purple), hourly average (–)SMR (orange), 1-min (–)SMR (green), 1-min (–)SYM-H (red), and hourly average (–)SYM-H index (pink) from Figure 3 on semi-log axis. Inset shows zoom for index values between 250 and 400 nT.

tightly clustered around the GPD fit (see Figures 3a–3e), they fall well within the 95% CI. The details of each GPD fit are different, as can be seen in Figures 3f–3h and 4. Comparing their appearance on linear axes, the fits to 1 hr indices are similar (Figure 3f). Despite scatter between the underlying empirical distribution, the GPD fits track each other for Dst [1-hr, 1957], hourly mean SYM-H, and hourly mean SMR indices. Similarly, the 1-min SYM-H and SMR indices (Figure 3g) show some scatter in the underlying empirical distribution but the GPD fits appear similar on a linear axis. Where the 1-hr Dst (from 1957) and 1-min SMR empirical distribution and best GPD fits are compared (Figure 3h), the distributions do not track each other. When the best fit GPD curves are displayed with a semi-log axis (Figure 4), the differences between the curves based on shape and scale parameters tabulated in Table 1 can be more clearly seen. Up to around 350 nT there is reasonable agreement between the 1-hr SYM-H, 1-hr SMR and Dst indices, beyond this point the GPD curves diverge and the 1-hr SYM-H and 1-hr SMR curves exhibit longer tails than either Dst sample. 1-minute SYM-H and SMR distributions track each other reasonably closely and show longer tailed behavior than any of the 1-hr indices up to around 650 nT. We note that Dst (from 1957) gives a bounded GPD distribution  $\xi < 0$  consistent with previous authors (Acero et al., 2018), as does 1-min SYM-H, whereas the other samples are unbounded  $\xi > 0$ . Whilst a bounded GPD distribution would have significant implications for extrapolating to the most extreme, Carrington-class events, we emphasize that the uncertainties contain both bounded and unbounded GPD fits. Our EVT based estimates should not be

extrapolated beyond the aforementioned estimated upper limit to the size of events, Dst  $\sim 2,000$ – $2,500$  nT (Liu et al., 2020; Vasyliūnas, 2011).

Risk is typically framed in terms of return level, for example, the 1 in 100 year return level is the amplitude that the index will exceed, on average, once in 100 years. If we rank-order the events in decreasing size, so that rank 1 is the largest observed event, then the rank is simply the number of times an event of that size has been seen to occur in the time interval of observation. The rank, divided by total number of events, is the survival function, so that our GPD fits translate directly into return level plots. Rather than invert the GPD fits, we use an expression for the return level that directly arises from the EVT (Coles, 2001, p. 103), as detailed in the SI. Estimates for the 5, 10, 50 and 100 year return levels for each index are tabulated in Table 2. Table 3 presents the percentage enhancement of return level estimates, relative to equivalent return period estimate and 95% CI for Dst (from 1957). Large confidence intervals for return levels translate to a large range in the confidence interval for percentage enhancement. Figures 5a–5e provide return level plots for the index timeseries. Return level estimates and

**Table 1**

Parameters of the Generalized Pareto Distribution Fit to Declustered Peaks-Over-Threshold Are Tabulated for Indices (–) Dst [1-hr, 1957], (–)Dst [1-hr, 1973], Hourly Average (–)SYM-H [1-hr, 1981], Hourly Average (–)SMR [1-hr, 1973], (–)SMR [1-min, 1973], and (–)SYM-H [1-min, 1981]

Index	Year	$u$	$n$	$\xi$	$\sigma$
Dst [1-hr, 1957]	1957	250	36	–0.13 [–0.39, 0.14]	88 [51, 125]
Dst [1-hr, 1973]	1972	250	21	0.01 [–0.41, 0.43]	73 [29, 117]
SYM-H [1-hr, 1981]	1981	250	22	0.06 [–0.35, 0.47]	80 [33, 126]
SMR [1-hr, 1973]	1972	250	26	0.10 [–0.28, 0.48]	70 [32, 107]
SYM-H [1-min, 1981]	1981	250	26	–0.06 [–0.39, 0.27]	109 [54, 164]
SMR [1-min, 1973]	1972	250	32	0.00 [–0.31, 0.32]	97 [52, 142]

Note. Year refers to the first year of the time series and  $n$  is the number of distinct events in the timeseries that is, declustered peaks over threshold. The GPD fit parameters consist of  $u$  the threshold,  $\xi$  the shape parameter and  $\sigma$  the scale parameter. Estimates of the 100-year return level for the indices based on the GPD model are presented. Estimates of the lower and upper bounds of 95% confidence intervals are in brackets.

**Table 2**

*Estimates for 5, 10, 50, and 100 Year Return Levels (RL) Based on the Generalized Pareto Distribution Model Are Presented for Each Index*

Index	5-year RL	10-year RL	50-year RL	100-year RL
Dst [1-hr, 1957]	337 [304, 371]	384 [342, 427]	485 [409, 560]	524 [424, 624]
Dst [1-hr, 1973]	315 [276, 353]	362 [310, 414]	476 [365, 587]	527 [369, 685]
SYM-H [1-hr, 1981]	331 [287, 375]	385 [320, 450]	537 [351, 723]	617 [328, 906]
SMR [1-hr, 1973]	324 [287, 361]	373 [318, 429]	516 [358, 673]	592 [348, 835]
SYM-H [1-min, 1981]	371 [319, 424]	433 [363, 504]	583 [426, 739]	650 [429, 870]
SMR [1-min, 1973]	359 [316, 403]	420 [352, 489]	603 [381, 824]	703 [352, 1054]

*Note.* Estimates of the lower and upper bounds of 95% confidence intervals are in brackets.

confidence bounds are overplotted with the rank-order estimates for the return levels of extreme events within the timeseries (the same events that form the basis of the GPD fits in Figure 3) and we can again see that the extreme events are, with the exception of the largest event, well within the 95% CI. The upper bound ( $\xi < 0$ ) of Dst (from 1957) is evident in the curved return level in Figure 5a. Comparing the return level estimates based on Dst (from 1957) to that for Dst (from 1973) in Figure 5b, 95% CI bounds are larger for the estimate based on the shorter time period but contain the estimate for the full available timeseries. Though SMR can have up to 120 stations and SYM-H is limited to 6 stations, comparing the 1-min SYM-H and SMR return level estimates (Figure 5f), the similarity between best GPD fits (Figure 3g) equates to return period estimates that are also consistent, despite the scatter between SMR and SYM-H POT empirical distributions. Despite the difference in numbers of stations, the extreme value analysis of the two 1-min timeseries returns very similar results. The hourly average SYM-H and SMR index return levels are compared to 1-hr Dst (Figures 5g and 5h). Comparing the return levels in Table 2 and the percentage enhancement over Dst (from 1957) in Table 3, it is seen that, up to 10 year return periods, the return level estimates are largely similar. At 50-year return periods and at 100-year return periods, the return level estimates are higher for 1-hr SYM-H and SMR than for those obtained for 1-hr Dst. The 1-min SYM-H and SMR index return levels are compared to 1-hr Dst (Figures 5i and 5j). The return level estimates for the 1-min indices are larger than those for Dst; the percentage increases are progressively bigger for longer return periods. Similarity between the 1-min SYM-H and SMR return period plots equates to similar percentage increases above the 1-hr Dst return level plot. From Tables 2 and 3, 1-min SYM-H and SMR return level estimates exceed those for Dst by increasing percentages for 5, 10, 50 and 100 year return periods.

SYM-H and Dst have been compared in the literature before, differences in quiet time and Sq variation subtraction are mentioned as possible sources of differences between the indices (Gannon & Love, 2011; Katus & Liemohn, 2013; Wanliss & Showalter, 2006). Wanliss and Showalter (2006) compared hourly-averaged SYM-H values to corresponding hourly Dst values. They found a linear relationship between Dst and SYM-H, with a break in the fit at  $-300$  nT; they found that increased activity coincided with larger differences between Dst and SYM-H, and that intense storms have deviations that are usually less than 20 nT. Wanliss and Showalter (2006) highlight multiple possible sources of differences between the indices, including spatial non-uniformity in ring current variations, magnetic field-aligned currents, and ionospheric currents at low latitudes; geologic

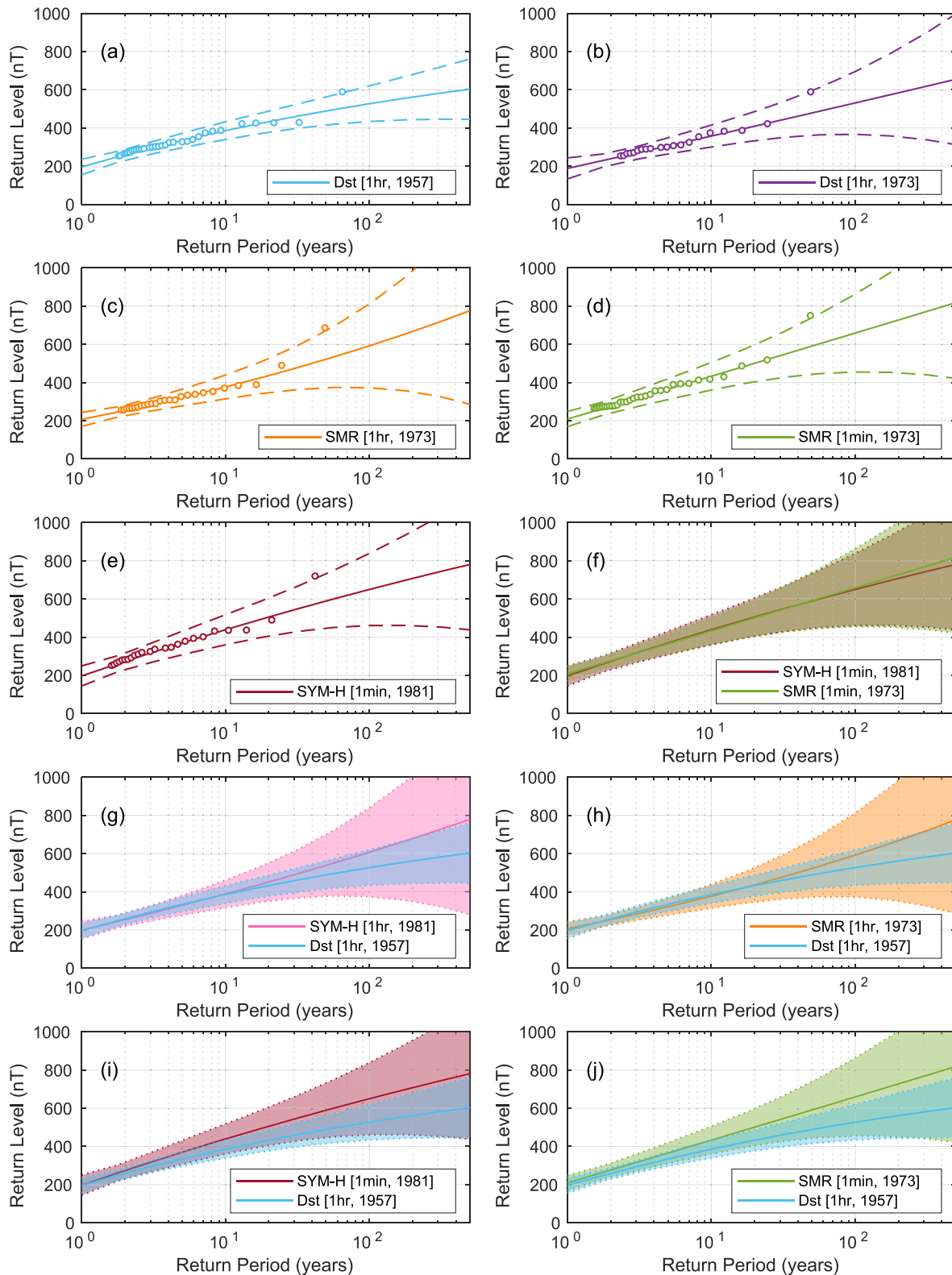
**Table 3**

*Percentage Enhancement of 5, 10, 50, and 100 Year Return Level Estimates (RL) Relative to the Estimates for the Dst [1-hr, 1957] Index Timeseries*

Index	5-year RL	10-year RL	50-year RL	100-year RL
Dst [1-hr, 1973]	-8% [-21%, 4%]	-7% [-22%, 8%]	-2% [-26%, 22%]	1% [-30%, 32%]
SYM-H [1-hr, 1981]	-2% [-17%, 14%]	1% [-18%, 19%]	10% [-23%, 43%]	15% [-29%, 59%]
SMR [1-hr, 1973]	-4% [-17%, 9%]	-2% [-18%, 14%]	7% [-24%, 37%]	12% [-29%, 54%]
SYM-H [1-min, 1981]	11% [-8%, 29%]	14% [-7%, 34%]	21% [-9%, 50%]	23% [-12%, 59%]
SMR [1-min, 1973]	9% [-7%, 26%]	12% [-6%, 31%]	21% [-9%, 51%]	25% [-14%, 64%]

*Note.* Based on results tabulated in Table 2, estimates of the 95% confidence intervals are in brackets.





**Figure 5.** Return level plots. (a) (–)Dst [1957–2021] (blue), (b) (–)Dst [1973–2021] (purple), (c) hourly average (–)SMR (orange), (d) 1-min (–)SMR (green), and (e) (–)SYM-H indices. For each timeseries, extreme value theory (EVT) return level estimates (solid line) with 95% confidence intervals (dashed line) are plotted. Circles indicate empirical estimates of the return levels for observed events (declustered exceedances of  $u = 250$  nT). The EVT return level estimates are compared between (f) 1-min (–)SMR and 1-min (–)SYM-H. The EVT return level estimates are compared between 1-hr (–)Dst [1957–2021] and (g) 1-hr (–)SYM-H (pink), (h) 1-hr (–)SMR, (i) 1-min (–)SYM-H, (j) 1-min (–)SMR indices.

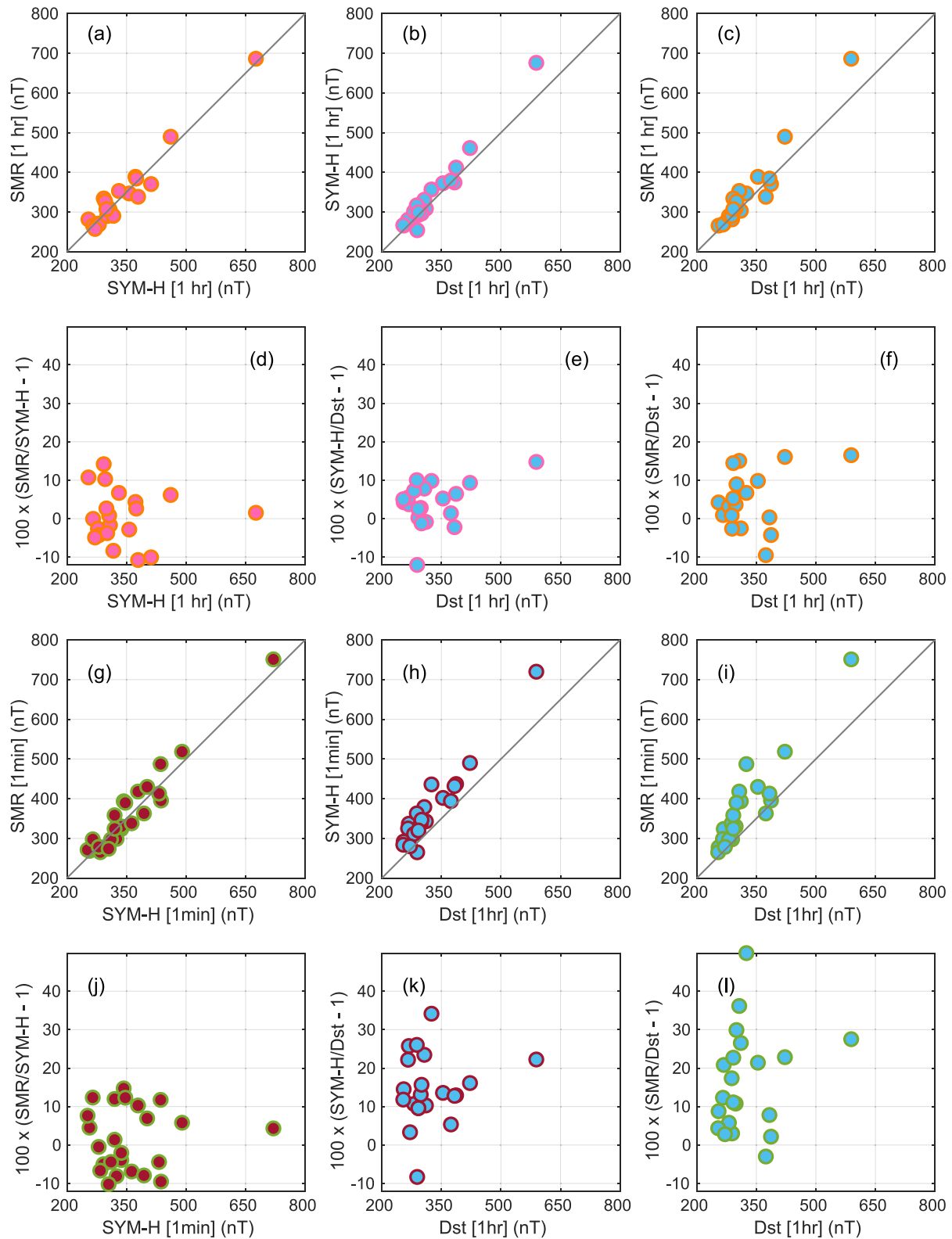
conductivity structures beneath the stations used to compute the indices are also variable; furthermore, transients such as those formed by the compression of the magnetopause by coherent structures in the solar wind result in fluctuations that vary on timescales shorter than the cadence of Dst observations. Katus and Liemohn (2013) also show that the differences between SYM-H and Dst are largest during super-storms. They find a difference of approximately 20% between the indices storm peaks. A record of Dst [1-min, 1985–2009] is studied in their analysis, the authors agree with a previous study (Gannon & Love, 2011) linking the differences between Dst [1-min, 1985–2009] and SYM-H indices to greater sensitivity of SYM-H to auroral zone ionospheric currents, as a result of the additional stations at higher magnetic latitudes in the SYM-H network.

In Figure 6, the POT values recorded by different indices are compared for events where both indices register a POT for the same event. Index record lengths vary so identified POT samples also vary from index to index; moreover, for some events one index will exceed 250 nT while another does not cross this threshold, thus a sub-sample of events are established for each pair of indices, where POT values have been identified in both of the index timeseries. Figure 7 gives an indication of which events will be included in the analysis of Figure 6; filled circles represent POT values of the SYM-H or SMR indices for which a corresponding POT has been identified in Dst, open circles indicate an POT value in the SYM-H or SMR timeseries for which no corresponding Dst POT has been identified. POT identification is consistent across the indices for extreme events with estimated return periods greater than 5 years. Comparing the POT values in Figure 6, we find that the SYM-H and SMR indices most closely track each (panels a,d); with percentage enhancement of  $\pm 10\%$  for 1-hr SMR over 1-hr SYM-H. 1-min SYM-H and 1-min SMR also show agreement (panels g, j), even for the largest event in the study with SYM-H  $\sim 750$  nT on 14 March 1989. SYM-H is based on only 6 stations, whereas SMR can include data from up to 120 stations. These results suggest that it is the temporal, rather than spatial, sampling that is the key driver of differences found between 1-min SMR and Dst. We find that the hourly average SYM-H and SMR POT values correspond to similar values of Dst (panels b,c,e,f). We find that the most significant difference between POT values arises when 1-hr Dst is compared to the 1-min SYM-H and SMR indices. There is greater deviation to the left of the diagonal (panels h and i), and percentage enhancement of either 1-min index over Dst ranges between  $-10\%$  and  $+50\%$  (panels k,l), with a weak positive trend seen for greater percentage enhancement for larger values of Dst. In their analysis, Gannon and Love (2011) name the inclusion of higher latitude stations in the SYM-H network as a possible source of difference between 1-min SYM-H and 1-min Dst. This comparison indicates that increased cadence of the SYM-H index relative to Dst has a greater impact than the inclusion of higher latitude stations. Figure 7 shows empirical estimates of the return levels for events identified in the timeseries. Comparing return level estimates for 1-hr indices (Figure 7a), we find that the rank-order empirical estimates for return periods of POT values are comparable, up to periods of 10 years. When the 1-min indices are compared to 1-hr Dst from 1957 (Figure 7b), the empirical return level plots do not track each other; in particular, the relationship between these indices is not linear and for extreme values these geomagnetic indices are not directly interchangeable.

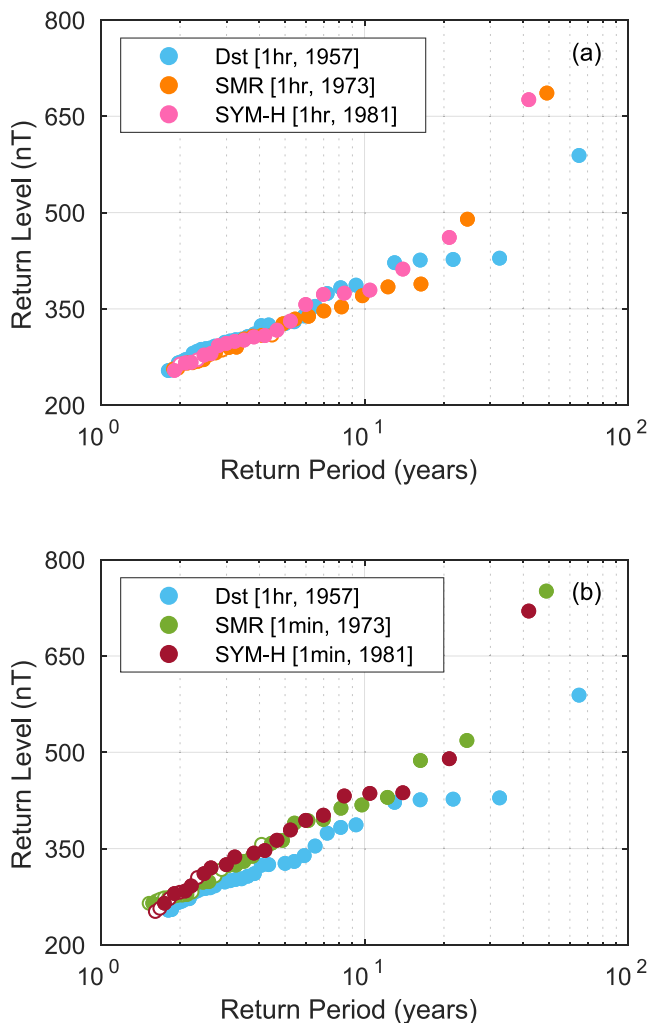
#### 4. Discussion and Conclusions

We performed a comparative extreme value analysis between the recent SuperMAG SMR index, the SYM-H index, and the Dst index, the latter being a benchmark for assessing the risk of geomagnetic storms. Compared to Dst, the SYM-H index offers greater temporal resolution and the SMR index has comparatively higher spatial and temporal resolution. As such, the SYM-H and SMR indices are expected to more closely resolve spatially and temporally localized extremes in the ground magnetic field disturbance (Newell & Gjerloev, 2012; Wanliss & Showalter, 2006). For extreme events, the contributions from the ring current (which these indices are designed to monitor) and contributions from magnetopause current (e.g., sudden commencement), substorm current wedge (both field-aligned currents and change in cross-tail current), or other magnetosphere-ionosphere current systems cannot be discriminated solely from the Dst, SYM-H or SMR indices (see e.g., Ohtani, 2022). Regardless of their source, such extremes are important since these short-time duration spikes in ground magnetic field perturbation are responsible for ground induced currents and hence damage to systems such as power transmission grids.

Whilst the time resolution of the indices does not vary within the record, the number of stations available to construct SMR increases with time (Bergin et al., 2020; Newell & Gjerloev, 2012). The spatial sensitivity of SMR is initially comparable with that of SYM-H but by halfway through the record has increased to over 50 stations. The derivation of the SMR index (see Section 2) involves averaging over stations within four local time sectors;



**Figure 6.** Comparison of peak-over-threshold (POT) events across indices. Corresponding POT values for individual events as parameterized by (a) 1-hr SYM-H and 1-hr SMR, (b) 1-hr Dst and 1-hr SYM-H, (c) 1-hr Dst and 1-hr SMR, (g) 1-min SYM-H and 1-min SMR, (h) 1-hr Dst and 1-min SYM-H, and (i) 1-hr Dst and 1-min SMR. Percentage enhancement of index relative to another for (d) 1-hr SMR over 1-hr SYM-H versus 1-hr SYM-H, (e) 1-hr SYM-H over Dst versus Dst, (f) 1-hr SMR over Dst versus Dst, (j) 1-min SMR over 1-min SYM-H versus 1-min SYM-H, (k) 1-min SYM-H over Dst versus Dst, and (l) 1-min SMR over Dst versus Dst.



**Figure 7.** Comparison of event return levels across indices. Empirical estimates of the return levels for peak-over-threshold (POT) values are compared for (–)Dst[1957–2021] (blue), 1-hr (–)SMR (orange), and 1-hr (–)SYM-H indices (pink) (top panel), and (–)Dst[1957–2021] (blue), 1-min (–)SMR (green), and 1-min (–)SYM-H indices (red) (bottom panel). Open circles indicate POT events that were identified in SMR or SYM-H but not Dst that is, the (–)Dst index did not exceed the 250 nT threshold during the event. Filled circles correspond to the events presented in Figure 6.

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#### Data Availability Statement

The Dst and SYM-H indices used in this paper were provided by the WDC for Geomagnetism, Kyoto (<http://wdc.kugi.kyoto-u.ac.jp/wdc/Sec3.html>) and were obtained from the WDC Kyoto interface (Dst: <https://wdc.kugi.kyoto-u.ac.jp/dstae/index.html>), (SYM-H: <https://wdc.kugi.kyoto-u.ac.jp/aeasy/index.html>). The SMR index was retrieved from the SuperMAG interface (<https://supermag.jhuapl.edu/indices/>). The download date was 6 December 2022.

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increasing spatial resolution within each local time sector will act to increase the sensitivity of the index to geomagnetic perturbations. The reasonably close (–10%,+15%) coincidence of 1-min SMR and SYM-H POT values (Figures 6g and 6j) suggests that the significantly different station numbers supporting the indices are not a strong driver of differences in the POT return levels. This suggests that the change in station numbers with time in SMR does not significantly affect our results.

From the GPD distributions fitted to the data, we found that for these more extreme events, SYM-H and SMR predict disturbances that are more intense than previously obtained using Dst, however, the increase does not increase linearly with event severity. Regarding comparison with previous estimates for Dst return levels, we would expect that fitting specific parametric distributions such as power laws or lognormals would also yield significant increases in the estimated return levels. This implies that selection of indices can have significant implications for risk assessment and planning mitigation against space weather hazard, for example, for 1-in-10-year return levels, our estimates based on 1-min SYM-H or SMR exceed those based on Dst by about 12%, while the 1-in-100-year return levels show enhancement of almost 25%.

Risk is commonly parameterized by estimated intensity of the 1 in 100 years event; our results suggest that a more nuanced approach to index selection may be required. Our results suggest that for more extreme events, these geomagnetic indices are not directly interchangeable, they contain different, complementary information and should be used together in any detailed analysis. Such an analysis may improve the correlation between return levels and the impact of severe geomagnetic storms.

In their comparison of Dst and SYM-H, Katus and Liemohn (2013) call attention to the role that global storm indices play in model assessment; the authors highlight that the interchangeable use of indices, without accounting for the differences between the indices, could bias data-model comparison tests. These authors recommended the use of multiple indices in determination of low-to middle-latitude magnetic perturbations and suggested that difference between indices could be considered as an error estimate on these values. Some recent studies such as Brenner et al. (2021) use the Dst, SYM-H, and SMR indices interchangeably. Our results show that the indices are not interchangeable, particularly during extreme events, and further emphasize the consideration that should be given to index choice in model validation and cross-study comparison.

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# Supporting Information for “Extreme event statistics in Dst, SYM-H and SMR geomagnetic indices.”

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**Contents**

1. Extreme Value Statistics: Details of the analysis technique employed in the main text are given, with reference to the requirement for an appropriate threshold selection.

2. Threshold stability analysis: Selection of an appropriate threshold for the analyses presented in the main text. The graphical analysis methods and results (Figure S1) are discussed.

3. Return level: The method of conversion of GPD fit parameters to an N-year return level is explained. Return level estimates are presented (Figure S2) that clarify the necessity of event declustering but they are supplementary to the findings presented in the main text.



## 1. Extreme Value Statistics

EVT provides a framework to model the probability of extreme events that are rarely observed in a timeseries and to extrapolate these models to unobserved levels. The basis for the POT method (Davison & Smith, 1990) is that, for a stationary process, peak exceedances over a suitably high threshold, separated in time by a suitable cluster interval, have a limiting distribution which asymptotically approaches the generalized Pareto distribution (GPD) whose survival function takes the form:

$$\bar{H}(y; \xi, \sigma) = \begin{cases} \left(1 + \frac{\xi y}{\sigma}\right)^{-1/\xi}, & \xi \neq 0 \\ \exp\left(-\frac{y}{\sigma}\right), & \xi = 0, \end{cases} \quad (1)$$

defined on  $\{y : y > 0 \text{ and } (1 + (\xi y)/\sigma) > 0\}$ . The GPD limiting distribution has two parameters: scale ( $\sigma$ ), and shape ( $\xi$ ). The three extreme value distributions in the generalized Pareto family are distinguished by shape parameter,  $\xi$ . If  $\xi < 0$  the distribution of excesses has an upper limit of  $u - \sigma/\xi$  but the distribution is unbounded if  $\xi = 0$  or if  $\xi > 0$ . For each geomagnetic index timeseries, the GPD model is fit to an empirical distribution comprised of peak-over-threshold values from each distinct event.

The specification of an appropriately high threshold is a requirement for the validity of the GPD model. Bias stems from thresholds that are too low due to violation of the asymptotic basis of the model. High variance occurs where thresholds are too high and few events with which the model can be estimated are identified. It is usual to deal with this bias-variance trade-off by identifying the lowest threshold at which the model provides a reasonable approximation. This ideal threshold is estimated in practice by testing for threshold stability, if the GPD is a valid model for peaks over a threshold  $u_0$ , then for all

higher thresholds  $u$ , the same GPD model should also apply (Pickands, 1975). GPDs fit to peaks over each threshold are expected to have identical shape parameter,  $\xi_{u_0} = \xi_u$  and scale parameters related by  $\sigma_u = \sigma_{u_0} + \xi(u - u_0)$ . The mean residual life (Hall & Wellner, 1981) or mean of the excesses of the threshold  $u$  is

$$E(X - u | X > u) = \frac{\sigma_u}{1 - \xi} = \frac{\sigma_{u_0} + \xi u}{1 - \xi}. \quad (2)$$

Accordingly, for  $u > u_0$ ,  $E(X - u | X > u)$  is a linear function of  $u$ . Based on (Coles, 2001, pp. 78 - 81), two graphical methods are used here to investigate threshold stability, (i)  $E(X - u | X > u)$  plotted against increasing  $u$  is the mean excess function plot where a region of linearity indicates the values of  $u$  for which the GPD model is stable (ii) plots of maximum likelihood estimates of  $\xi$  and  $\sigma^*$  parameters as a function of threshold, where  $\sigma^* = \sigma_u - \xi u$  is the reparameterized scale parameter, and estimates of  $\xi$  and  $\sigma^*$  are expected to be constant above  $u_0$ .

## 2. Threshold stability analysis

In order to determine a suitable threshold at which to carry out the EVT analysis, the mean excess function is plotted in Figure S1a. The mean excess function is estimated by

$$E(X - u | X > u) = \frac{1}{n_u} \sum_{i=1}^{n_u} (x_i - u) \quad (3)$$

where  $x_1, \dots, x_{n_u}$  are the  $n_u$  observations that exceed  $u$ . 95% confidence intervals (CI) are calculated from standard error of the mean. The four lines diverge above 250 nT and have approximately linear behaviour with a slope closer to zero for Dst, and positive slopes for SMR. 1 minute SMR has the largest positive gradient. The reparameterized scale and shape parameter estimates are plotted for increasing thresholds in Figures S1b-e. Declustering is applied to the threshold exceedances as described in the main body of the

paper and the MATLAB EVIM gpd function is used to generate a maximum likelihood estimate (MLE) with 95% CI for the GPD parameters. With few events in the declustered observation tails, there is variability in the estimates for  $\xi$  and particularly  $\sigma^*$ , particularly at high thresholds but in general the threshold stability plots are in agreement with other Dst EVT studies; 250 nT (Acero et al., 2018) and 280 nT (Tsubouchi & Omura, 2007). We select a threshold of  $u = 250$  nT throughout. Above the specified thresholds, with declustering applied, the following threshold exceedances are identified: 36 exceedances of 250 nT in Dst [1957], 21 exceedances of 250 nT in Dst [1973], 26 exceedances of 250 nT in hourly mean SMR and 32 exceedances in 1-min SMR indices.

### 3. Return Level

Information about the number of threshold exceedances per distinct event is retained through the extremal index,  $\theta$ , (Leadbetter, 1983) which acts as measure of the clustering in the timeseries (Moloney et al., 2019). For a timeseries with  $n_u$  exceedances of the threshold  $u$ , and  $n_e$  distinct events identified using the POT procedure detailed above, the runs estimator for the extremal index (Smith & Weissman, 1994) is  $\theta = n_e/n_u$ , that is the inverse of the mean number of threshold exceedances in each distinct event.

The  $N$  year return level,  $z_N$  is the threshold that the geomagnetic index is likely to exceed, on average, once every  $N$  years. Informally following (Coles, 2001, p. 81), if the sequence of  $n_e$  observed peaks-over-threshold identified in the timeseries,  $Y_1, \dots, Y_{n_e}$  for  $n_e$  distinct events, are samples of common distribution  $Y$ , and the GPD is an appropriate model for  $Y$ ,  $\bar{H}(y; \xi, \sigma)$  is the probability that any peak-over-threshold will take a value greater than  $y$ ;  $\bar{H}(y) = Pr\{Y > y\}$ . Generalizing peaks-over-threshold model to the timeseries as a whole then, let  $\zeta_u$  be the probability of an individual observation exceeding the

threshold  $u$  and  $\theta$  be the inverse of the mean number of threshold exceedances in each distinct event represented by the peak-over-threshold. The probability of a given observation exceeding the threshold  $u$ , constituting a peak over the threshold  $u$ , and exceeding the level  $x$  is then the product  $\zeta_u \theta \bar{H}(x - u)$ . The level  $x_m$  that is exceeded once every  $m$  observations can be estimated by solving  $\theta \zeta_u \bar{H}(x_m - u) = m^{-1}$ . The level that is likely to be exceeded once every  $N$  years is found by substituting  $Nn_y$  for  $m$  where  $n_y$  is the number of observations per year. Thus the GPD survival function,  $\bar{H}$ , forms the basis for the  $N$  year return level estimate,

$$z_N = \begin{cases} u + \frac{\sigma}{\xi} \left[ (Nn_y \zeta_u \theta)^\xi - 1 \right], & \xi \neq 0 \\ u + \sigma \log(Nn_y \zeta_u \theta), & \xi = 0. \end{cases} \quad (4)$$

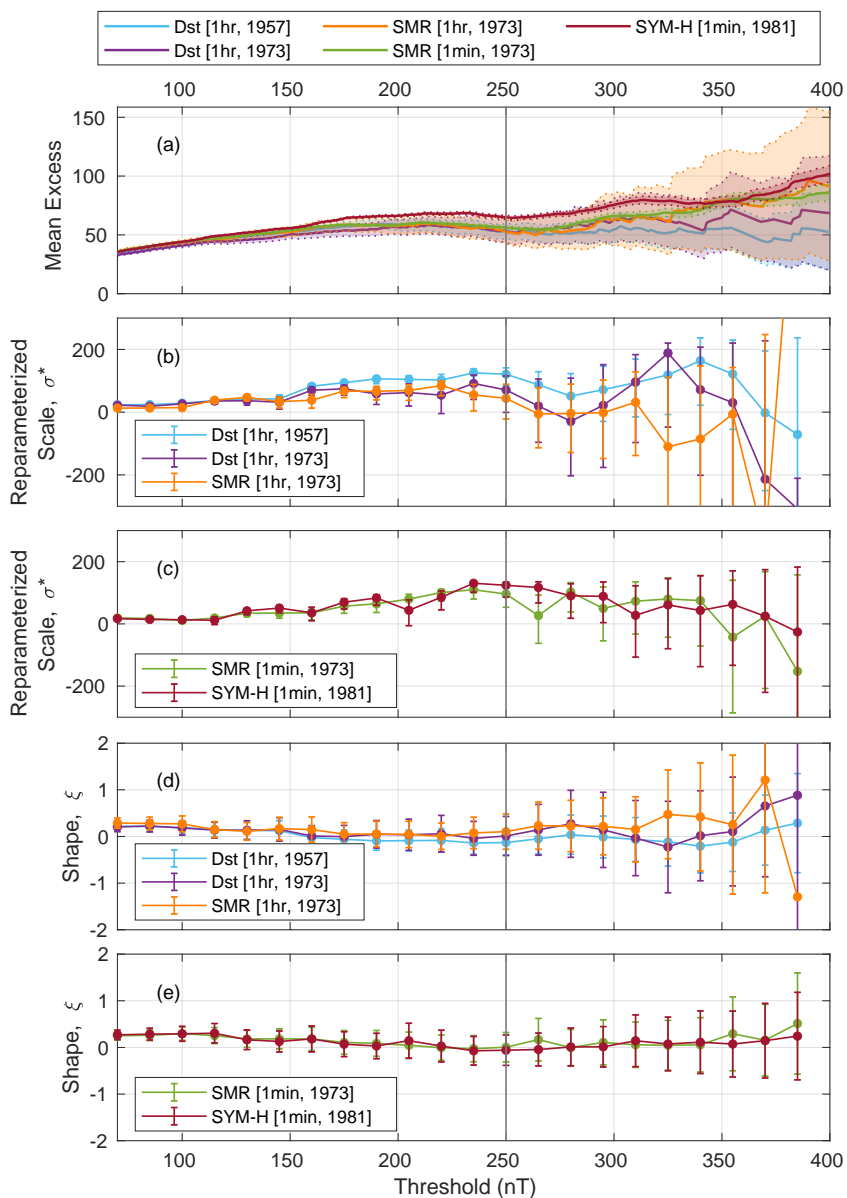
The estimated return levels ( $z_N$ ), that is the thresholds that the indices are expected to exceed, on average, once every  $N$  years, are compared in Figure S2. The return level estimate (Equation 4) combines the parameters of the GPD models with an estimate for  $\zeta_u$  (the probability of an individual observation exceeding the threshold  $u$ ). Empirical estimates for the observation return levels can be calculated. If the  $n$  exceedances of the threshold  $u$  are arranged in descending order,  $Y_1 > \dots > Y_k > \dots > Y_n$ , so that  $Y_1$  is the largest threshold exceedance and  $Y_n$  is the smallest threshold exceedance, taking  $N$  to be the total duration of the index observation period, the empirical estimate for the return levels of each threshold exceedance,  $Y_k$ , is  $N/k$ . Figure S2 compares the empirical return level estimates for the total set of threshold exceedances with the empirical estimates for return level of the cluster peaks within the timeseries. We see again that the declustering procedure is particularly relevant in an index such as 1-min SMR, where the cadence of the observation is far less than the typical duration of an excursion in the observed timeseries. Overplotted on the empirical threshold exceedance and cluster peak return

levels is the GPD estimate for the return level (Eq. 4) as a function of return period with the estimated 95% confidence interval calculated using the delta method (Coles, 2001, p. 82). The return level estimate is higher for 1-min (-)SMR than the hourly (-)SMR index, and both (-)SMR estimates are higher than the (-)Dst index return levels. The upper bound ( $\xi < 0$ ) of (-)Dst[1957] is evident in the curved return level in Figure S2a. The return level estimate for 1-min (-)SMR exceeds the upper uncertainty bound of the (-)Dst[1957] return level estimate. This result points towards a systematic under-estimation of the 100-year return level in the Dst index, as discussed in the main text.

## References

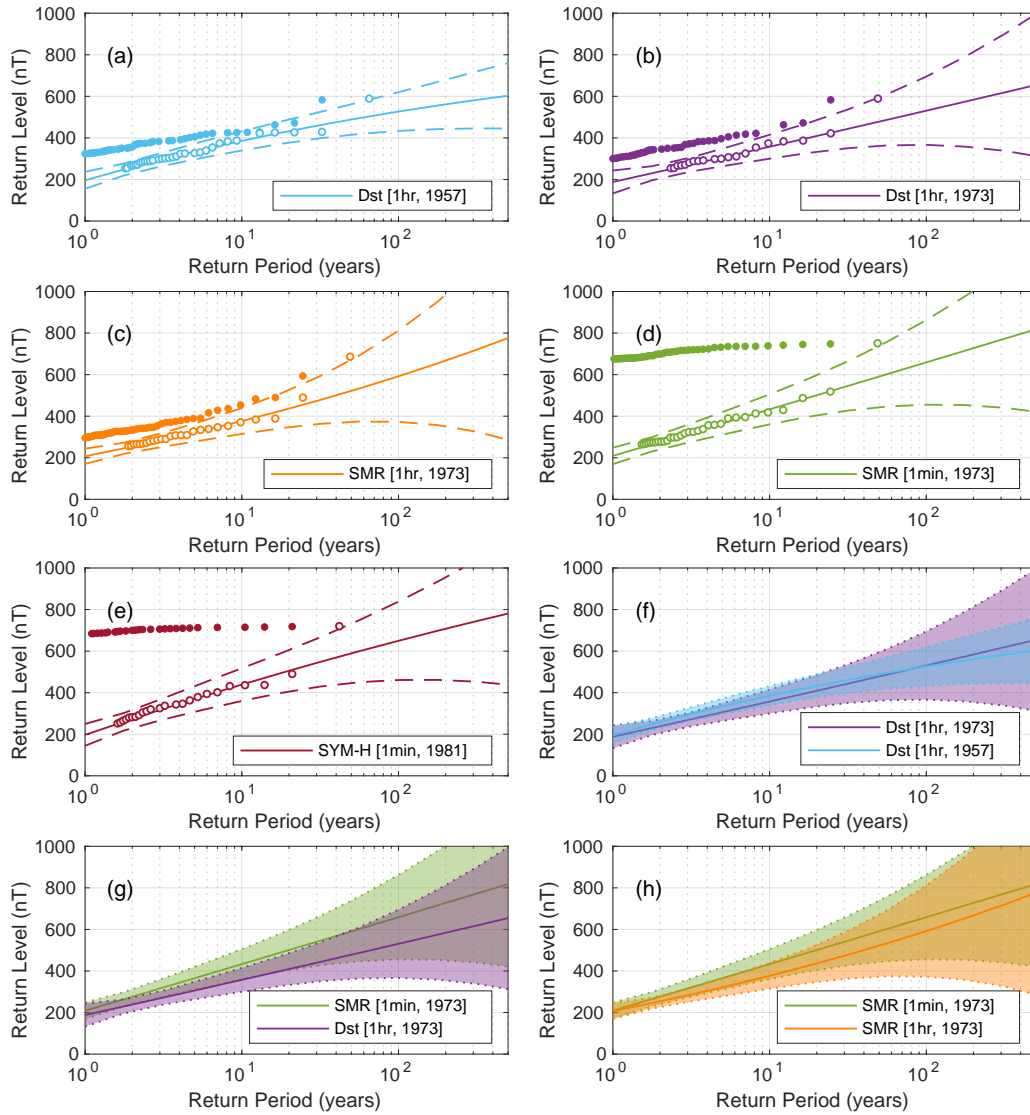
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**Figure S1. Threshold stability analysis.** For (-)Dst [1957 - 2021] (blue), (-)Dst [1973 - 2021] (purple), hourly average (-)SMR (orange), 1-min (-)SMR (green), and (-)SYM-H (red) indices. (a) Mean excess plot (solid) with approximate 95% confidence intervals (shaded). (b,c) Estimate of reparameterised GPD scale parameter,  $\sigma^*$ , for declustered timeseries as a function of threshold, with 95% confidence interval from MLE of GPD fit indicated by error bars. (d,e) Estimate of GPD shape parameter,  $\xi$ , for declustered timeseries, plotted versus threshold with 95% confidence interval indicated by error bars.

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**Figure S2. Return level plots.** for (a) (-)Dst [1957 - 2021] (blue), (b) (-)Dst [1973 - 2021] (purple), (c) hourly average (-)SMR (orange), (d) 1-min (-)SMR (green), and (e) (-)SYM-H (red) indices. For each timeseries, return level estimates (solid line) with 95% confidence intervals (dashed line) are plotted. Scatter points indicate empirical estimates of the return levels for sets of exceedances of  $u = 250$  nT, where all exceedances are retained (filled) and where exceedances are declustered over a cluster interval of 48 hrs (open). The return level estimates are compared for (f) (-)Dst[1957 - 2021] (blue) and (-)Dst[1973 - 2021] (purple), (g) (-)Dst[1973 - 2021] and 1-min (-)SMR (green), and (h) hourly average (-)SMR (orange) and 1-min (-)SMR (green).